

# **Accelerating Drug Development Through the Use of Combination Phase II/III Designs**

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# Outline of Presentation

- Adaptive design introduction
- Potential uses of Phase II/III combination designs
- Example of **2** look Phase II/III design
  - Type I error control
  - Power assessment
  - Timing of Look #1
- Example of **3** look Phase II/III design
- Comparison vs. separate Phase II & Phase III
- Future Work
- Conclusions

# Some Types of Adaptive Designs

	<u>Incorporated here</u>
1. Stop for futility	Yes
2. Stop early for efficacy	Yes
3. Sample size re-estimation	Future
4. Add in new arms	No
5. Change randomization ratio	No
6. Change primary endpoint(s)	No
7. Change test statistic	No

This talk describes combination Phase II/III designs incorporating 1-2

# Situations Where Adaptive Designs are Most Useful

The following also applies to interim analysis in general:

1. Follow-up short relative to duration of enrollment
2. Short time from LPV to Interim Analysis decision
3. Randomization via IVR if arms can be dropped

Note: If follow-up is long but onset of action is rapid can often overcome #1 by using result on primary endpoint (or surrogate) but from early visit

# Phase II/III Combination Trial - Situations Where this could be Useful

- Range of doses will likely cover optimal dose
  - so that arms will not need to be added
- At most 5 doses still under consideration
- Major safety concerns not likely to apply to very many doses
  - otherwise separate Phase II is probably preferable

# Phase II/III Combination Trial - Situations Where this could be Useful, cont'd

- Certain Single Study Submissions
  - Fast Track (if Phase II, III endpoints are the same)
  - New stage of disease, or closely related disease
  - New patient population
  - New combination therapy
  - Orphan indication or other rare disease

## Phase II/III Combination Trial - Situations Where this could be Useful, cont'd

- Where sponsor would otherwise carry out Phase II, Phase III #1, Phase III #2 in sequence due to limited funds
- Methodology may also sometimes be useful in place of a multi-armed Phase III trial
  - e.g., in place of past Phase III trial that included 5 doses (50-fold range) and Placebo

# Aims of Phase II/III Combination Studies

## Considered Here:

- Combine dose selection and confirmatory stages
  - start with 2-5 doses + placebo
- Not stop for success in first 50% of study
  - due to safety database needs
- Stop for success as early as possible once we have enough patients for safety
- Stop study early if all doses are ineffective

# Phase II/III Combination Design

## - Two Look Case

1. At start randomize to PL,  $D_1, \dots, D_k$
2. At Interim
  - Choose "best" dose  $D_{BD}$
  - *Decision rule specified in protocol & administered by independent group*
  - Stop if all  $D_i$  futile
  - Randomize to PL,  $D_{BD}$  from now onwards
3. At Final
  - Test  $D_{BD}$  vs. PL at level  $\alpha_2$
  - Efficacy demonstrated if test statistic  $\geq Z_{1-\alpha_2}$

## Two Look Case, cont'd

- Suppose data is normally distributed
- $Z_{iD_j}$  = test statistic at look #i for  $D_j$  vs. PL
- Test at 1-sided  $\alpha_i$  for  $i=2$  (onwards)
- $< Z_{1-\alpha_0}$  is futility decision rule used to stop study at look #1 (with corresp. CP)
- $n_1$  per group at look #1,  $n_2$  extra per group at look #2
  - also generalized to allow unequal #s per group

# Two Look Case- Type I Error

Type I error is given by

$$\sum_{j=1}^k P( Z_{1D_j} \geq Z_{1-\alpha_0} \cap Z_{2D_j} \geq Z_{1-\alpha_2} \cap D_j \text{ selected} )$$

Note: this applies whatever decision rule is used to select  $D_j$

## Two Look Case - Type I Error, cont'd

Suppose decision rule at look #1 is to choose  $D_j$  corresponding to highest  $Z_{1D_j}$

Type 1 error can be shown to be given by

$$\int_{Z_{1-\alpha_0}}^{\infty} \Phi \left( \sqrt{\frac{n_1}{n_2}} v - Z_{1-\alpha_2} \sqrt{\frac{n_1 + n_2}{n_2}} \right)$$

$$* P \{ \max (Z_{1D_1} \dots Z_{1D_k} = v) \} dv$$

See also Simon et al (1994), Hsu et al (1997), Todd & Stallard (2001)

# Calculation of Critical Alpha levels

- Equate previous equation to  $\alpha = 0.025$  (1-sided)
- 2d Integral evaluated numerically making use of results from Dunnett (1955)
- For given  $n_1, n_2, k, \alpha$ , solve for  $\alpha_2$
- Could increase  $\alpha_2$  even further (as Tsong et al, 1997) by allow for unspent T1E resulting from stopping study due to futility
  - For now, not made use of this in case sponsor decides to override DSMB recommendation to stop and continues with the two-arm trial for stage 2

# Critical Alpha levels

<b>k</b>	<b>10%</b>	<b>20%</b>	<b>30%</b>	<b>50%</b>	<b>99.9%</b>	<b>Dunnett's alpha</b>
1	0.02500	0.02500	0.02500	0.02500	0.02500	0.02500
2	0.01919	0.01751	0.01645	0.01510	0.01350	0.01348
3	0.01667	0.01443	0.01306	0.01136	0.00944	0.00941
4	0.01517	0.01266	0.01115	0.00933	0.00733	0.00731
5	0.01414	0.01147	0.00990	0.00803	0.00603	0.00601

# Determination of Power

- This is analytically more complex and so for now is determined by simulation (100,000 runs for each example)
- Example with  $k=4$  active arms
  - $\mu/\sigma = (0, 0.07, 0.14, 0.21, 0.22)$
  - $\mu/\sigma = (0, 0.06, 0.12, 0.18, 0.24)$
  - $\mu/\sigma = (0, 0.03, 0.11, 0.20, 0.23)$
  - $\mu/\sigma = (0, 0.08, 0.16, 0.24, 0.16)$
- # patients in trial fixed at  $1500 = 5n_1 + 2n_2$

# Power and Timing of Interim #1

$n_1$	$n_2$	Interim #1	$\alpha_2$	Power			
				Case 1	Case 2	Case 3	Case 4
30	675	4.3%	0.01788	85.8%	82.4%	82.9%	84.8%
60	600	9.1%	0.01550	88.4%	86.0%	87.4%	<b>87.2%</b>
90	525	14.6%	0.01381	<b>89.4%</b>	<b>87.8%</b>	<b>89.4%</b>	<b>88.1%</b>
120	450	21.1%	0.01247	<b>89.8%</b>	<b>88.7%</b>	<b>89.9%</b>	<b>88.3%</b>
150	375	28.6%	0.01133	<b>89.3%</b>	<b>88.6%</b>	<b>89.6%</b>	<b>87.9%</b>
240	150	61.5%	0.00865	83.3%	83.9%	84.2%	83.7%

## Timing of Interim #1 with k=4

- When determining timing of Interim #1 seek to balance
  - (a) Need for high power in current study
  - (b) High chance that "best dose" is selected at interim #1
  - (c) Time of interim is late enough for good dose-response information to be obtained
- In examples with k=4, having interim at 20% approximately, gave highest power.
  - 20%-30% may be preferable when allow for (b)-(c)

## Extension to 3 or more Looks

1. At start randomize to PL,  $D_1, \dots, D_k$
2. At Interim #1
  - Choose "best" dose  $D_{BD}$
  - Stop if all  $D_i$  futile
  - Randomize to PL,  $D_{BD}$  from now onwards
3. At Interim # $i$  ( $i > 1$ ) and Final (look # $r$ )
  - Test  $D_{BD}$  vs PL at level  $\alpha_i$
  - Stop trial for efficacy if test statistic  $\geq Z_{1-\alpha_i}$

# Three Look Case - Type I Error

Type I error in normal case:

$$\sum_{j=1}^k P( Z_{1Dj} \geq Z_{1-\alpha_0} \cap Z_{2Dj} \geq Z_{1-\alpha_2} \cap D_j \text{ selected} )$$
$$+ \sum_{j=1}^k P( Z_{1Dj} \geq Z_{1-\alpha_0} \cap Z_{2Dj} < Z_{1-\alpha_2} \cap Z_{3Dj} \geq Z_{1-\alpha_3} \cap D_j \text{ selected} )$$

Note: this expression applies whatever decision rule is used to select  $D_j$

## Three Look Case - Type I Error, cont'd

- Suppose, as before, decision rule at look #1 is to choose  $D_j$  with highest  $Z_{1D_j}$
- Type 1 error can be expressed as a multivariate normal probability, evaluated numerically
- Dimensionality can be reduced by allowing for independent increments (Todd & Stallard, 2001)
- For given  $n_1, \dots, n_r, k, \alpha$ , and spending function (relating  $\alpha_2, \alpha_3, \dots, \alpha_r$ ) solve for  $\alpha_r$

# Three Look Case - Example

Suppose that:

- Pocock-like  $\alpha$ -spending function is used, i.e.,  
 $Z_{1-\alpha_2} = Z_{1-\alpha_3}$
- No look for early efficacy at Interim #1
  - may want to incorporate extreme Haybittle-Peto like bound at interim #1, using  $Z_{1-\alpha_1} = 6.0$
- Looks for efficacy at 75%, 100%
- Example values of  $\mu/\sigma$  as before, with  $k=4$
- # patients in trial fixed at  $1500=5n_1+2n_2+2n_3$

# Power and Timing of Interim #1

## - 3 Look Case

Interim #1	$\alpha_2 = \alpha_3$	Overall Power			
		Case 1	Case 2	Case 3	Case 4
4.0%	0.01159	84.5%	80.9%	81.6%	83.5%
9.1%	0.00974	87.3%	85.0%	86.3%	85.8%
13.5%	0.00878	<b>88.2%</b>	86.4%	<b>87.9%</b>	86.4%
20.1%	0.00777	<b>88.4%</b>	<b>87.5%</b>	<b>88.4%</b>	<b>87.1%</b>
25.9%	0.00714	<b>87.8%</b>	<b>87.5%</b>	<b>88.5%</b>	<b>86.7%</b>
32.5%	0.00659	86.9%	<b>87.1%</b>	<b>87.7%</b>	86.4%
52.4%	0.00552	83.2%	83.9%	83.9%	83.3%

# Sample Size Calculations for k=4

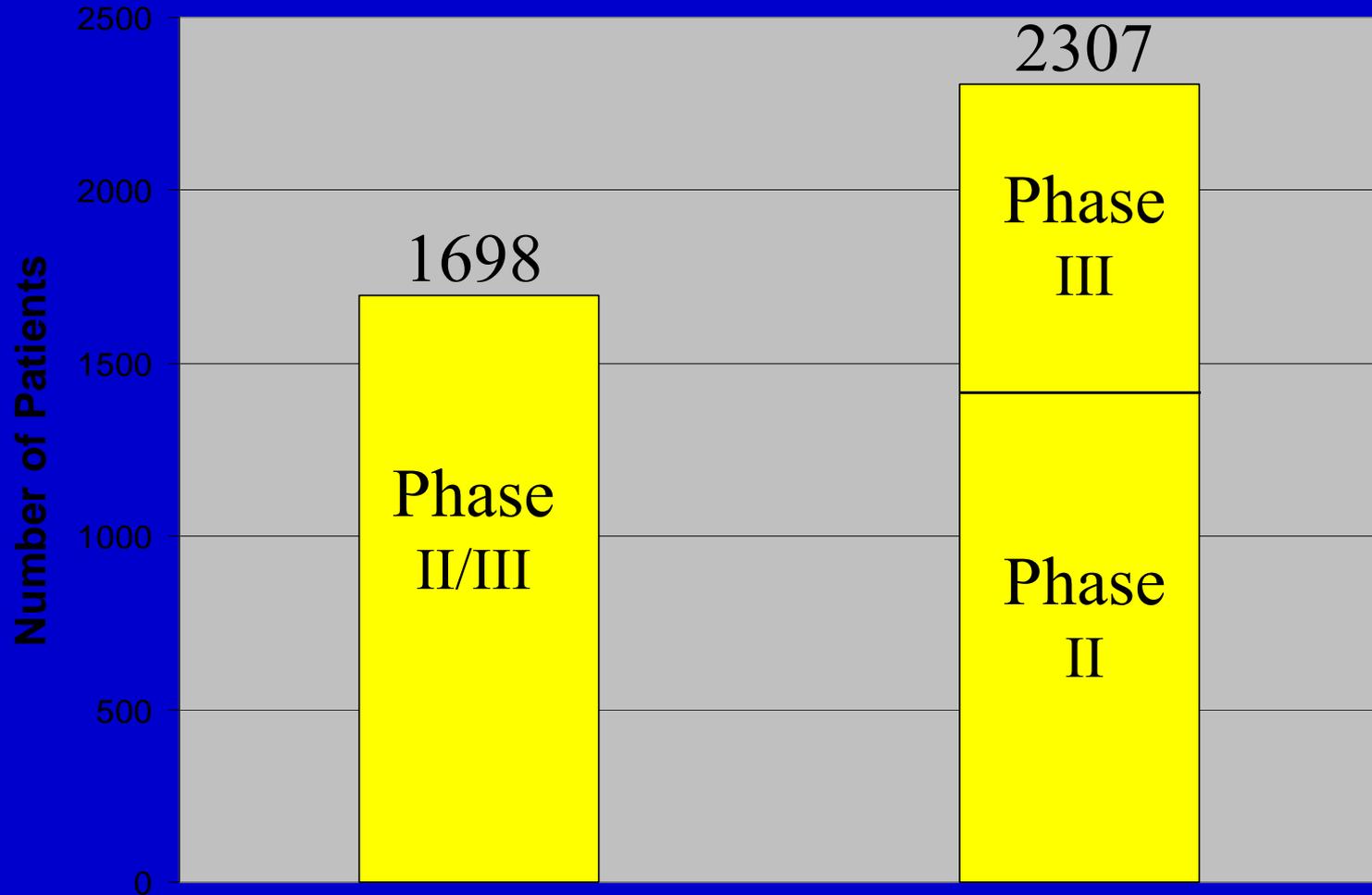
## - Separate Phase II & Phase III Trials

- Assumptions
  - Interim #1 at 20% in Phase II/III, i.e.,  $n_2 = 4 n_1$
  - Equal replication
  - $\mu/\sigma = (0, 0.07, 0.14, 0.21, 0.22)$
- For separate trials we require 2307 patients
  - 287 per arm (1435 total) for 5 arm Phase II based on Williams' test with 80% power
  - 436 per arm (872 total) for 2 arm Phase III based on  $\Delta/\sigma = 0.22$ , 90% power

# Sample Size Calculations for k=4 - Combined Phase II/III Trial

- Phase II/III (90% power) requires 1698 patients
  - $n_1=118$  in each of 5 groups prior to Interim #1
  - $n_2=472$  extra in PL and selected dose group
  - $1534 = 2 n_2 + 5 n_1$
  - Also have 164 patients enrolled between last interim patient and implementation of dropping of their arm
    - Assumes 12m enrollment (33 per week), primary endpoint at 30 days, 4w from LPV to drop 3 arms

# Relative Number of Patients



# Comparison of Timelines

## Phase II/III Study



## Phase II



## Phase III



# Phase II/III Combination Trial vs. Separate Phase II & Phase III

Benefits of Phase II/III trial in this example:

- **609 Fewer Patients** (2307 - 1698) needed as use Phase II data in final analysis
- **7.3 Months Saved**
  - 18.5 weeks due to enroll 609 fewer patients (33/w)
  - >3 months between LPI study #1 and FPI study #2

Assuming 30d duration in Phase II + 4w from LPO to dose selection  
+ 4w to start up 2nd study & gear up enrollment again

## Further Work

- More extensive evaluation of current Phase II/III design approach
  - $k = 2, 3, 5$  & broader sets of  $\Delta/\sigma$ , etc.
  - further comparison vs separate Phase II & Phase III
- Modified decision rules
  - Allow for shape of dose-response curve in decision rule
  - Extend calculations to take account of chance that  $D_j$  has safety problems, where dose with maximal  $Z_{1D_j}$  may not then be chosen
    - T1E still controlled
    - Impact on power needs assessment

## Further Work, cont'd

- Allow two "best" doses (& PL) to be kept after interim #1
- Incorporate sample size re-estimation
  - unblinded, extending Liu & Chi (2001), Cui et al (1999), or Bauer & Kohne (1994)
  - blinded, extending Gould & Shih (1991, 1998)

# Conclusions

- Methodology presented allows combination of dose-finding and confirmatory stage within one study
- Type 1 error is controlled exactly at 0.025
- Allows early stopping if all doses are clearly ineffective
- Can be combined with any alpha-spending function to enable stopping as early as possible, subject to meeting safety database needs

## Conclusions, cont'd

- Allows for dose selection based on Phase III primary endpoint
- Preliminary results show that if we start with as many as 4 doses of test drug, then having dose selection at 20% - 30% gives
  - high power
  - low chance of continuing with a sub-optimal dose
  - adequate dose-response information

## Conclusions, cont'd

- Preliminary results indicate that, where this approach is consistent with needs of program, it can:
  - Cut Drug Development Time  
**7.3 Months Time Saving** in Example
  - Cut Costs by Reducing Number of Patients  
**609 Fewer Patients Needed** in Example